



# Graph Construction: An Empirical Investigation on Setting the Range of the Y-Axis

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Graphs are an effective and compelling way to present scientific results. With few rigid guidelines, researchers have many degrees-of-freedom regarding graph construction. One such choice is the range of the y-axis. A range set just beyond the data will bias readers to see all effects as big. Conversely, a range set to the full range of options will bias readers to see all effects as small. Researchers should maximize congruence between visual size of an effect and the actual size of the effect. In the experiments presented here, participants viewed graphs with the y-axis set to the minimum range required for all the data to be visible, the full range from 0 to 100, and a range of approximately 1.5 standard deviations. The results showed that participants' sensitivity to the effect depicted in the graph was better when the y-axis range was between one to two standard deviations than with either the minimum range or the full range. In addition, bias was also smaller with the standardized axis range than the minimum or full axis ranges. To achieve congruency in scientific fields for which effects are standardized, the y-axis range should be no less than 1 standard deviations, and aim to be at least 1.5 standard deviations.

*Keywords:* Graph Design, Effect size, Sensitivity, Bias

One way to lie with statistics is to set the range of the y-axis to form a misleading impression of the data. A range set too narrow will exaggerate a small effect and can even make a non-significant trend appear to be a substantial effect (Pandey, Rall, Satterthwaite, Nov, & Bertini, 2015). Yet the default setting of many statistical and graphing software packages automatically sets the range as narrow as the data will allow. The problem of creating misleading graphs persists even when the full range is shown instead. As shown in the studies reported below, a

range set too wide also creates a misleading impression of the data by making effects seem smaller than they are. Here, I argue that for scientific fields that use standardized effect sizes and adopt Cohen's convention that an effect of  $d = 0.8$  is big, the range of the y-axis should be approximately 1.5 standard deviations (SDs).

How should the y-axis range of a graph be determined? Graph construction should account for the visual experience of the people reading the graphs (Cleveland & McGill, 1985; Kosslyn, 1994; Tufte, 2001) and the strong link between perception and cognition (Barsalou, 1999; Glenberg, Witt, & Metcalfe, 2013). When the visual size of the effect aligns with the actual size of the effect, the person reading the graph does not have to exert mental effort to decode effect size from the graph. Instead, the size of the effect is processed automatically. This increases graph fluency by making it easier to understand

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Data, scripts, and supplementary materials available at [osf.io/hw2ac](https://osf.io/hw2ac).

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Table 1. Overview of the five experiments.

Experiment	N	Effect sizes	Graph Type	Standardized condition <sup>1</sup>
1	9	0.1, 0.3, 0.5, 0.8	Bar graph	2 SDs
2	14	0.1, 0.3, 0.5, 0.8	Bar graph	1.4 SDs
3	13	0, 0.3, 0.5, 0.8	Bar graph with error bars	1.2 SDs
4	20	0, 0.3, 0.5, 0.8	Line graph	1.4 SDs
5	15	0, 0.3, 0.5, 0.8	Line graph	1 SD

Notes. <sup>1</sup>This refers to the range depicted in the standardized condition, so a range of 1.4 SDs is when the graph was centered on the grand mean and extended 0.7 SDs in either direction.

that an effect is big when it looks big and an effect is small when it looks small.

To increase graph fluency, the range of the y-axis should be selected to maximize compatibility between visual size and actual effect size (Kosslyn, 1994; Pandey et al., 2015; Tufte, 2001). However, the current literature fails to provide clear guidelines on how to achieve this compatibility. For example, some recommend displaying only the relevant range so that the axis goes from just below the lowest data point to just above the highest data point (Kosslyn, 1994). This would not achieve the recommended compatibility because small effects would look big. Others assert that the y-axis should always start from 0, particularly for bar graphs (Few, 2012; Pandey et al., 2015; Wong, 2010). This too could fail to achieve compatibility by making effects look too small.

In the case of scientific fields for which effect size is standardized based on standard deviation, the range of the y-axis should be a function of the standard deviation (SD). In behavioral sciences such as psychology and economics, for example, the mean effect size is approximately half a SD (Bosco, Aguinis, Singh, Field, & Pierce, 2015; Open Science Collaboration, 2015; Paterson, Harms, Steel, & Crede, 2016), and a standardized effect size of  $d = .8$  is considered a big effect (Cohen, 1988). Consequently, an appropriate range for the y-axis would be one to two SDs, which would be plotted as the group mean  $\pm 0.75$  SD (or  $\pm 0.5 - 1$  SDs). With this range, big effects such as a Cohen's  $d$  of  $.8$  would look big and small effects of  $d = .3$  would look small. In other words, this range would help achieve compatibility between the visual impression of the size of the effect and the actual size of the effect.

## Empirical Studies

The effect of visual-conceptual size compatibility on graph fluency was empirically tested in 57 participants across 5 experiments (see Table 1). The participants were naïve college students, which serves as an appropriate sample given that scientific results should be accessible and comprehensible to this population and not just to experts in one's field.

The stimuli were bar or line graphs that had been constructed from simulated data. Data were simulated from two (hypothetical) groups of participants by sampling from normal distributions in R (R Core Team, 2017). For one group, the data were drawn from a normal distribution with a mean of 50 and a standard deviation of 10 (as in a memory experiment with mean performance of 50% and SD of 10%). For the other group, the data were drawn from a normal distribution with a standard deviation of 10 and the mean at 49, 47, 45, or 42. These means correspond to effect sizes of  $d = 0.1, 0.3, 0.5,$  and  $0.8$ , respectively. In Experiments 3-5, the mean of 49 ( $d = 0.1$ ) was replaced with the mean of 50 ( $d = 0$ ). In Experiments 2-5, the data were re-sampled if the attained effect size differed by more than 0.1 from the intended effect size. Data were simulated 10 times for each of the four effect sizes to create 40 sets of data for each Experiment. In Experiments 1-3, the means of the simulated data were displayed as a bar graph depicting two groups of participants who engaged in different study strategies (spaced versus massed; see Figure 1). In Experiments 4-5, the means were used to determine the end points of a line graph, and the x-axis was labeled as "hours spent studying". For each set of data, three graphs were constructed that varied in the range of the y-axis. The full condition showed the full range from 0 to 100 on a hypothetical memory test. The minimal condition showed the smallest range necessary to see the data. The standardized condition was centered on the group

mean and extended by one to two SDs in either direction (the exact value differed across experiments, see Table 1 or the Appendix). Figure 1 shows several examples of graphs that served as stimuli. In Experiment 3, error bars were also included and explained to the participants. Within an experiment, the same set of 120 graphs (3 axis ranges x 4 effect sizes x 10 sets) were shown to the participants. Graphs were shown one at a time, order was randomized, and participants completed 4 blocks of 120 trials. In all experiments, the participants' task was to indicate whether there was no effect, a small effect, a medium effect, or a big effect for each graph by pressing 1, 2, 3 or 4 on the keyboard.

Graph fluency was measured using linear regressions rather than accuracy because regression coefficients have the advantage that they provide two separate measures. The slope provides an estimate of sensitivity to the magnitude of the effect depicted in the plot. A steeper slope indicates better sensitivity to effect size than a shallower slope. The intercept provides an estimate of bias. Two graphs could lead to similar levels of sensitivity but different levels of bias. Separate linear regressions were calculated for each participant for each y-axis range condition (full, standardized, and minimal).

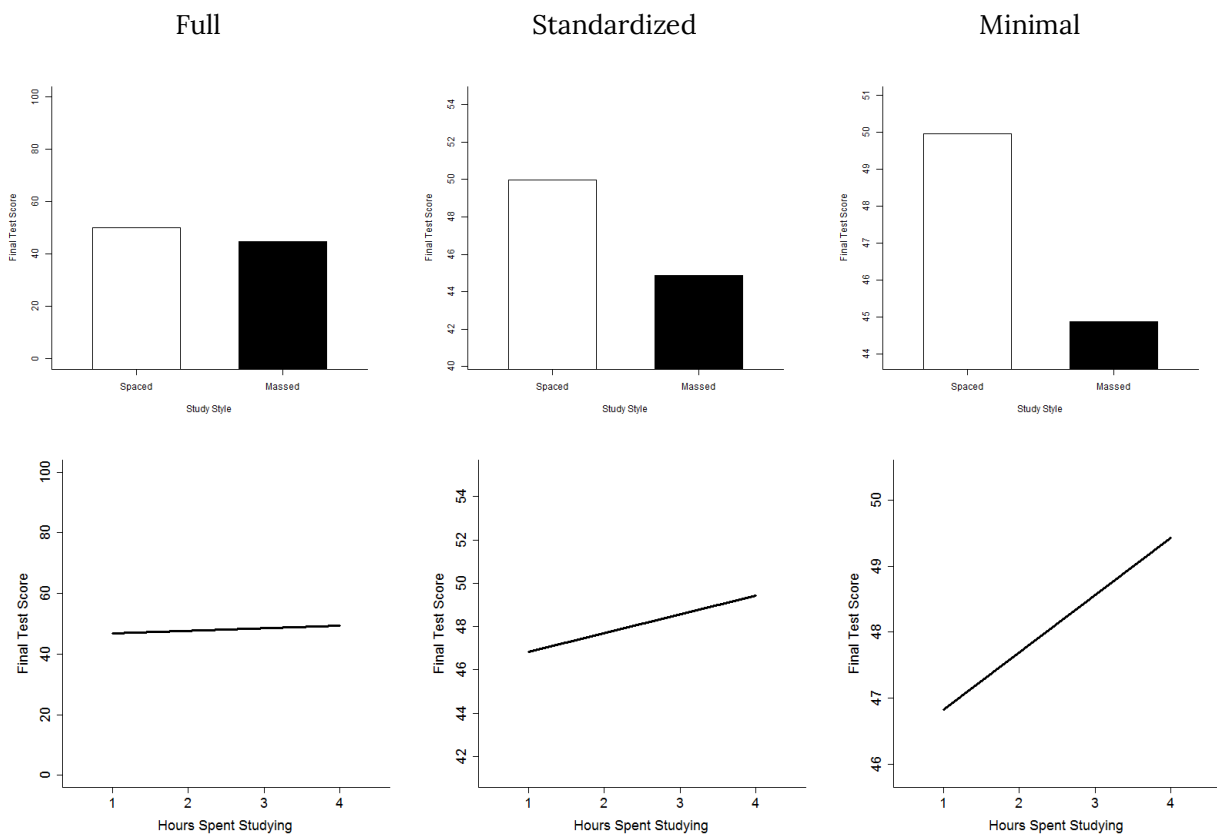


Figure 1. Sample stimuli in the experiments on bar graphs and on line graph. The bar graphs show final test score as a function of whether study style was spaced or massed. The line graphs show final test score as a function of hours spent studying from 1 to 4. Within each experiment, the same data were plotted using the full range from 0-100, the standardized range (in this case, the group mean +/- 0.7 SD), or the minimal range necessary to see the data. In this example, a medium effect (Cohen's  $d = 0.5$ ) was simulated for the bar graphs (top row) and a small effect (Cohen's  $d = 0.3$ ) was simulated for the line graphs (bottom row). The participant's task was to indicate whether there was no effect, a small effect, a medium effect, or a big effect.

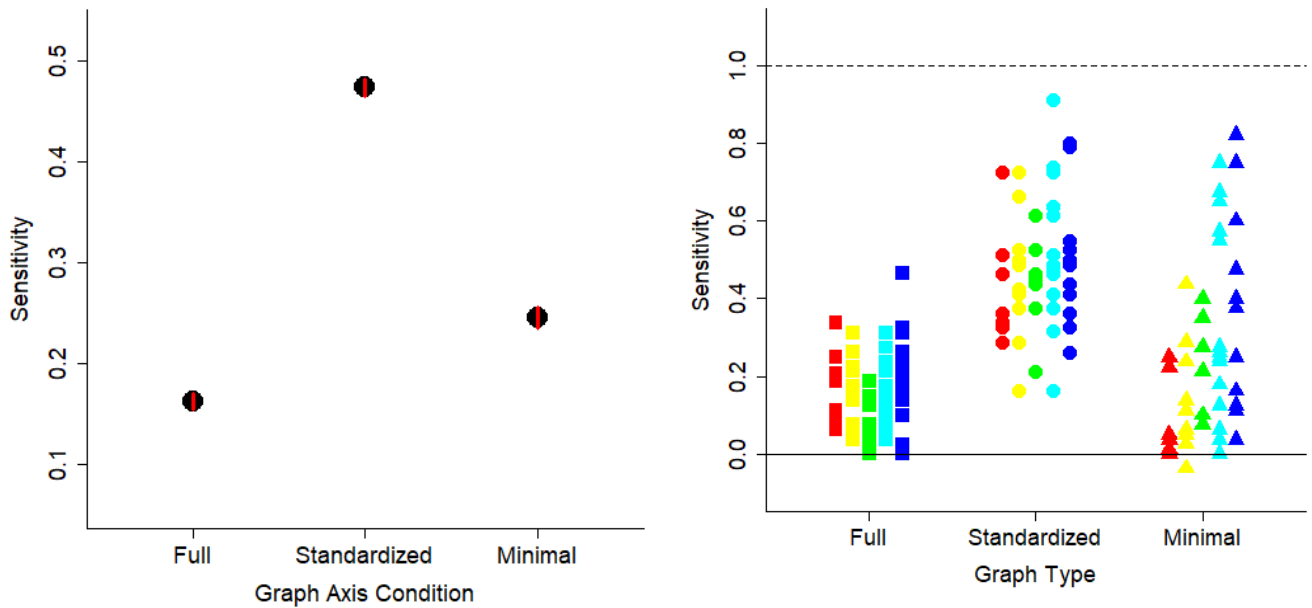


Figure 2. Sensitivity is plotted as a function of graph axis condition for the three types of graphs across all 5 experiments. Sensitivity was measured as the coefficient for the slope from regressions of actual effect size on estimated effect size. Only trials for which the graph depicted an effect size greater than  $d = 0.1$  are included (see supplementary materials for all the data). A higher sensitivity score corresponds to better performance, and a coefficient of 1 corresponds to perfect performance. A coefficient of 0 indicates chance performance. In the left panel, mean sensitivity across all experiments is shown. Error bars are 1 SEM calculated within-subjects, and are approximately the same size as the symbols. The right panel shows sensitivity for each participant for each experiment. The data are color-coded by experiment (e.g. red = Experiment 1, orange = Experiment 2) and are also laterally positioned from left to right within graph type category. Each point corresponds to one participant, and each participant has one symbol for each of the three graph types. The solid horizontal line at 0 shows the point of no sensitivity and the dashed horizontal line at 1 shows the point of perfect sensitivity.

In each regression, the dependent measure was response (on the scale of 1 to 4). The effect sizes were recoded to also be on a scale from 1 to 4 then centered by subtracting 2.5 so that perfect performance would produce a regression coefficient for the slope of 1 and an intercept of 2.5.

Figure 2 shows the mean slope coefficients across all 5 experiments. Sensitivity was best for the standardized graphs and worse for the full range graphs. Participants were better able to assess the size of the effect depicted in the graph for the standardized graphs, than for the minimal or full graphs. Participants were also less biased when viewing the standardized graphs. Figure 3 shows the mean bias across all 5 experiments. Bias scores were calculated as a percent bias based on the coefficients for

the intercept. A negative score indicates a bias to respond that effects were small, and a positive score indicates a bias to respond that the effects were big. For the full graphs, there was a large bias to respond that the effects were small. When looking at graphs with the full range, participants responded that almost all effects (86%) were null or small. For the minimal graphs, there was a large bias to respond that the effects were substantial. When looking at graphs with the minimal range for Cohen's  $d = 0.10 - 0.80$ , participants responded that the effect was big on 49% of the trials. In contrast, there was much less bias with the standardized graphs (see Supplemental Materials).

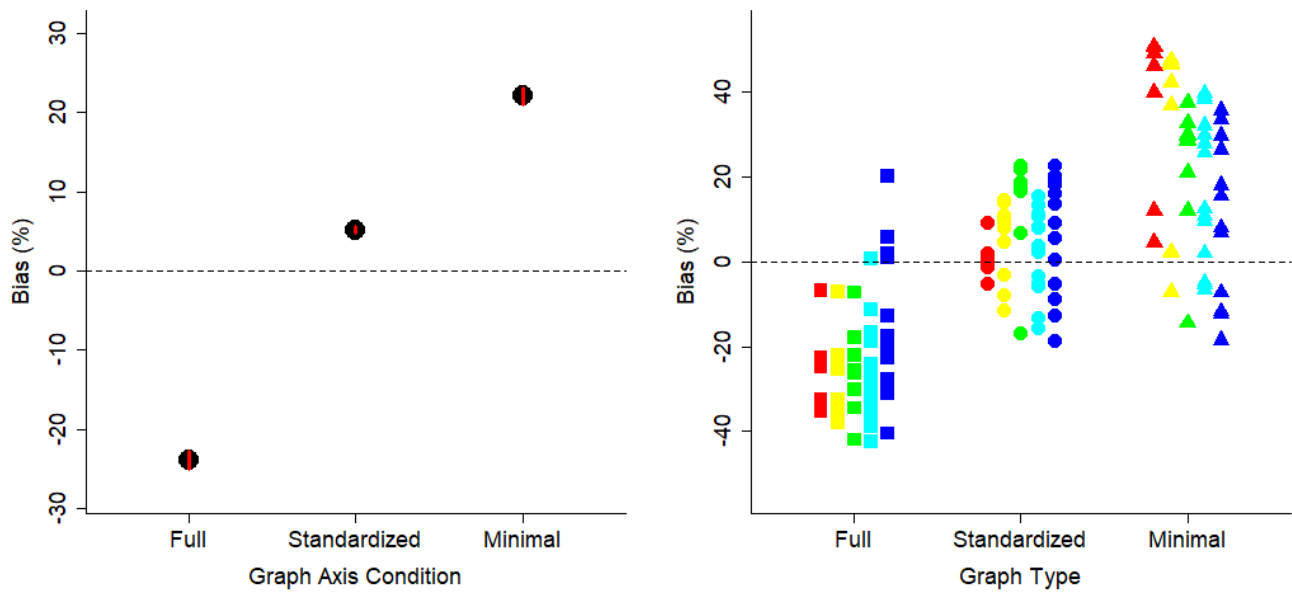


Figure 3. Bias (as a percentage) is plotted as a function of graph axis condition for the three types of graphs across all 5 experiments. A negative bias corresponds to responding that effects are smaller than they are, and a positive bias corresponds to responding that effects are bigger than their actual size. In the left panel, mean bias across all experiments is shown. Error bars are 1 SEM calculated within-subjects, and are approximately the same size as the symbols. The y-axis range is 4 SD. The right panel shows bias for each participant for each experiment. The data are color-coded by experiment (e.g. red = Experiment 1, orange = Experiment 2) and are also ordered from left to right within graph type category. Each point corresponds to one participant, and each participant has one symbol for each of the three graph types.

### Discussion

The visual impression of the size of an effect has a strong influence on the judged size of an effect. When the visual impression was compatible with the actual effect size, judgments of effect size were better calibrated and less biased compared with the typical default setting of showing the minimum range to display the data and the setting of showing the full potential range. Based on the current studies, the recommendation is to center the y-axis on the grand mean and extend the range 0.75 SDs in either direction so that the range of the y-axis is 1.5 SDs.

The current studies show improved sensitivity to effect size and reduced bias in estimating effect size when the range of the y-axis was centered on the grand mean of the data and extended approximately 0.7 SDs in either direction. The various studies used slightly different extensions ranging from 0.5 SDs to 1 SD. There were not large detectable differences in sensitivity or bias depending on the exact range that

was used, so the precise value of the y-axis range might not be critical. Rather, the key feature is that the visual size aligns with the actual size of the effect. The specific range to be used might vary as a function of the size of the error bars (the range should be large enough to encompass them), the size of the effect (the range would have to be extended for particularly large effects, such as was done with the current results), if doing so would make the range include nonsensical numbers (such as negative numbers for performance), and to achieve a consistent scale across multiple graphs to enhance across-graph comparisons. Given that the exact range in terms of SD could vary from plot to plot, it could be useful to indicate the range in SD units in the figure caption. This indication would be particularly useful in cases for which researchers do not include error bars.

The current experiments explored graphs of stimulated data from between-subjects designs. The recommendations likely generalize to within-subject designs with the caveat that the y-axis

should be a function of the denominator used to calculate the within-subjects effect size. For example, the denominator for Cohen's  $d_z$  is the square root of the sum of the squares of the standard deviations minus the product of the standard deviations and the correlation between the two measures. Graphs plotting within-subjects data could be  $\pm 0.75$  times this denominator (or one of the other suggested measures for within-subjects effects sizes; e.g. Lakens, 2013). In cases for which there are both between-subjects and within-subjects factors, the researchers will have to decide which denominator to use for the range depending on which effect they most want to emphasize.

It is debatable whether the recommendation offered here should be employed with bar graphs. Some have shown that graphs that start at a position other than 0 are deceptive (e.g., Pandey et al., 2015). The idea is that bar graphs should always start at 0 because the height of the bar signifies the value of the condition being represented. When the y-axis starts at a value greater than 0, the height of the bar corresponds to the difference between the condition's value and the starting point, rather than the condition's value itself. Consider the following example: imagine that group A scored 70% on a memory test and group B scored 60%. On a plot for which the y-axis starts at 50%, group A's score would appear twice as big as group B's score, even though they only scored 10% higher. The issue at hand concerns the visual impression of the data. If the graph gives the impression that the differences are big, and that aligns with the size of the effect, the graph would produce compatibility between vision and true effect size. If, however, the impression is that one group's performance was twice as good as the other group's performance, this would produce a misleading impression of the data. The current experiments cannot speak to which impression was experienced because participants were asked to rate the size of the effect as being no effect, small, medium, or big, rather than quantifying the size of one bar relative to another. The specific task used here did not permit measuring the spontaneous impression given by the graphs. One option is for researchers to use alternative types of graphs to avoid the issue. Alternatives include point graphs and a newly-designed type of graph called a hat graph (Witt, 2019).

The recommendation to set the y-axis range to be 1.5 SDs does not generalize to fields for which the SD is unknown or irrelevant for interpreting effect size. For these fields, previous recommendations such as Tufte's Lie Detector Ratio could be appropriate (Tufte, 2001). But for scientific fields that rely on standard deviation to interpret effect size, this is the first empirically-based recommendation that provides clear guidelines for constructing graphs to communicate the magnitude of the effects.

Maximizing compatibility between visual size and conceptual size improved comprehension of the effects shown in the graphs. The data presented in the graphs were exactly the same, yet participants were less biased and were more sensitive to the size of the depicted effect when the axis range was one to two SDs. Furthermore, emphasizing SD and effect size in graph construction could help shift researchers' focus to effect size, rather than statistical significance. Indeed, effect size (as measured with Cohen's  $d$ ) provides a better measure for discriminating real effects from null effects than  $p$  values or Bayes factors (Witt, 2019). Such a shift could help guard against practices that have contributed to recent failures to replicate in various scientific fields (Camerer et al., 2016; Open Science Collaboration, 2015).

In his famous book on how to lie with statistics, Huff noted that as long as the y-axis is correctly labeled, "nothing has been falsified – except the impression that it gives" (Huff, 1954, p. 62). The impression matters. Researchers should select the range of the y-axis so that small effects look small and big effects look big (based on the field's adopted conventions). A simple way to do this is to set the range to be 1.5 (or more) standard deviations of the dependent measure. That this improves graph comprehension is both intuitive and is now supported by empirical evidence.

### Open Science Practices



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### Author contribution

Witt is solely responsible for this manuscript. The author read and approved the final manuscript.

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### Conflict of Interest Statement

The author declares there were no conflicts of interest.

### References

- Barsalou, L. W. (1999). Perceptions of perceptual symbols. *Behavioral and Brain Sciences*, 22, 577-660.
- Belia, S., Fidler, F., Williams, J., & Cumming, G. (2005). Researchers misunderstand confidence intervals and standard error bars. *Psychological Methods*, 10(4), 389-396. doi: 10.1037/1082-989X.10.4.389
- Bosco, F. A., Aguinis, H., Singh, K., Field, J. G., & Pierce, C. A. (2015). Correlational effect size benchmarks. *Journal of Applied Psychology*, 100(2), 431-449. doi: 10.1037/a0038047
- Camerer, C. F., Dreber, A., Forsell, E., Ho, T.-H., Huber, J., Johannesson, M., . . . Chan, T. (2016). Evaluating replicability of laboratory experiments in economics. *Science*, 351(6280), 1433-1436.
- Cleveland, W. S., & McGill, R. (1985). Graphical Perception and Graphical Methods for Analyzing Scientific Data. *Science*, 229(4716), 828-833. doi: 10.1126/science.229.4716.828
- Cohen, J. (1988). *Statistical Power Analyses for the Behavioral Sciences*. New York, NY: Routledge Academic.
- Collaboration, O. S. (2015). Estimating the reproducibility of psychological science. *Science*, 349(6251), aac4716. doi: 10.1126/science.aac4716
- Cumming, G., & Finch, S. (2005). Inference by eye: confidence intervals and how to read pictures of data. *American Psychologist*, 60(2), 170-180. doi: 10.1037/0003-066X.60.2.170
- Few, S. (2012). *Show Me the Numbers: Designing Tables and Graphs to Enlighten* (Second Edition ed.). Burlingame, CA: Analytics Press.
- Glenberg, A. M., Witt, J. K., & Metcalfe, J. (2013). From revolution to embodiment: 25 years of cognitive psychology. *Perspectives on Psychological Science*, 8(5), 574-586.
- Huff, D. (1954). *How to Lie with Statistics*. New York, NY: W. W. Norton & Company.
- Kosslyn, S. M. (1994). *Elements of Graph Design*. New York: W. H. Freeman and Company.
- Lakens, D. (2013). Calculating and reporting effect sizes to facilitate cumulative science: A practical primer for t-tests and ANOVAs. *Frontiers in Psychology*, 4, 863. doi: doi:10.3389/fpsyg.2013.00863
- Morey, R. D., Rouder, J. N., & Jamil, T. (2014). *BayesFactor: Computation of Bayes factors for common designs (Version 0.9.8)*, from <http://CRAN.R-project.org/package=BayesFactor>
- Pandey, A. V., Rall, K., Satterthwaite, M. L., Nov, O., & Bertini, E. (2015). *How Deceptive are Deceptive Visualizations?: An Empirical Analysis of Common Distortion Techniques*. Paper presented at the Proceedings of the 33rd Annual ACM Conference on Human Factors in Computing Systems, Seoul, Republic of Korea.
- Paterson, T. A., Harms, P. D., Steel, P., & Crede, M. (2016). An assessment of the magnitude of effect sizes: Evidence from 30 years of meta-analysis in management. *Journal of Leadership & Organizational Studies*, 23(1), 66-81.
- Revelle, W. (2018). *psych: Procedures for Psychological, Psychometric, and Personality Research*. Retrieved from <https://cran.r-project.org/package=psych>

- R Core Team. (2017). *R: A language and environment for statistical computing*. Retrieved from <https://www.r-project.org>
- Tufte, E. R. (2001). *The Visual Display of Quantitative Information* (Second Edition ed.). Cheshire, CT: Graphics Press.
- Witt, J. K. (2019). *Introducing hat graphs*. Retrieved from [psyarxiv.com/sg37q](https://psyarxiv.com/sg37q).
- Witt, J. K. (2019). Insights into criteria for statistical significance from signal detection analysis. *Meta-Psychology*, 3, MP.2018.871. doi: 10.15626/MP.2018.871
- Wong, D. M. (2010). *The Wall Street Journal Guide to Information Graphics: The Dos & Don'ts of Presenting Data, Facts, and Figures*. New York, NY: W. W. Norton.



## Appendix: Experimental Details

### Experiment 1: Bar Graphs with Axis Range of 2 SD

Participants judged the size of effects depicted in bar graphs that were constructed with three axis range options.

#### Method

**Participants.** Nine students in an introductory psychology course participated in exchange for course credit. In this and all subsequent experiments, the number of participants was maximized within a pre-determined time limit.

**Stimuli and Apparatus.** Graphs were constructed in R (R Core Team, 2017). For each graph, two means were generated. One mean was 50, and the other mean was 49, 47, 45, or 42. These equated to effect sizes of Cohen's  $d = .1, .3, .5,$  and  $.8,$  respectively. To add some noise to each graph, each mean was drawn from a normal distribution centered on the desired mean with 1000 samples and a standard deviation of 10. The means were presented in bar graphs (see Figure A1). The left bar was white and labeled "Spaced" and the right bar was black and labeled "Massed". For each set of simulated data, three bar graphs were constructed that corresponded to the three y-axis range conditions. For

the full graphs, the y-axis range went from 0 to 100. For the minimal graphs, the y-axis went from the smallest data value minus 1 to the largest data value plus 1. For the standardized graphs, the mean of the two groups was calculated, and 1 SD (10) was added in either direction to set the y-axis range. This process of creating 3 graphs for each set of data was repeated 10 times for each of the 4 effect sizes for a total of 120 graphs. Graphs were 500 pixels by 500 pixels and were shown on a 19" computer monitors with 1028 x 1024 resolution.

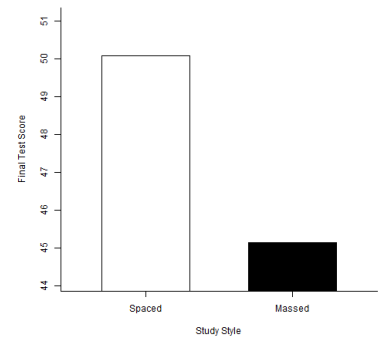
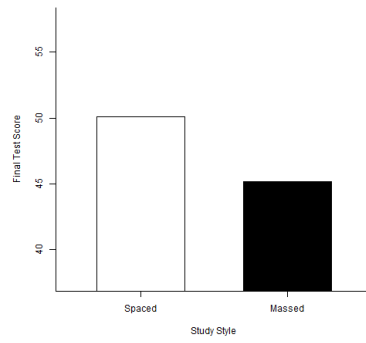
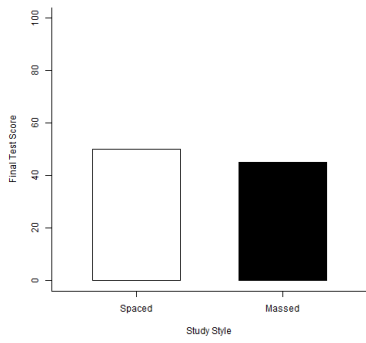
**Procedure.** After providing informed consent, each participant was seated at a computer. They were given the following instructions: "You will see graphs showing the effect of study style on final test performance. There were two study styles. Massed is like cramming everything at once at just before the exam. Spaced refers to studying a little bit every day for weeks before the exam. The y-axis shows final test performance, with higher value meaning better performance."

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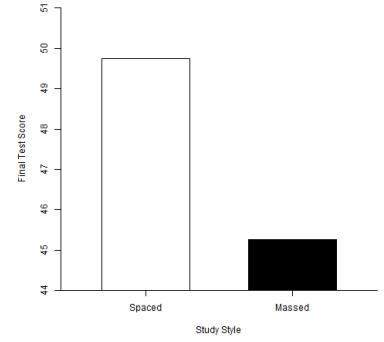
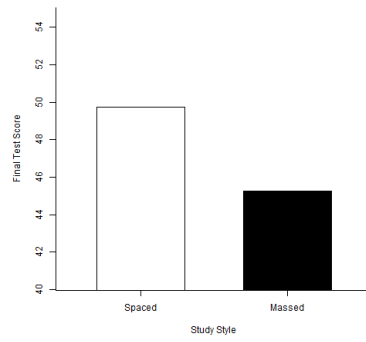
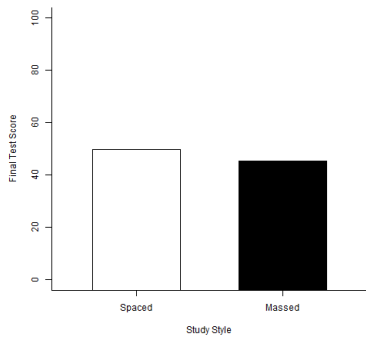
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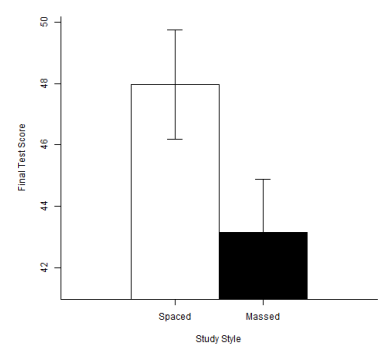
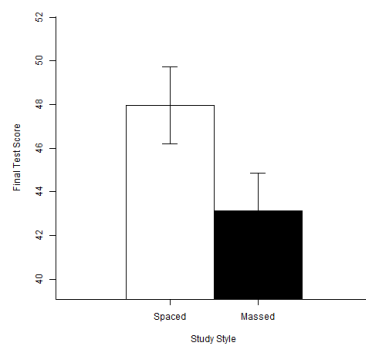
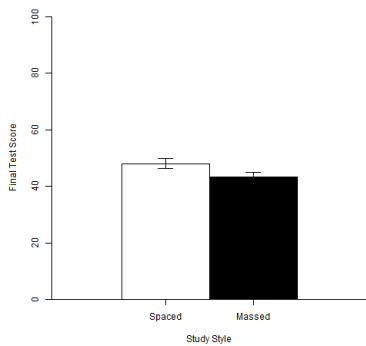
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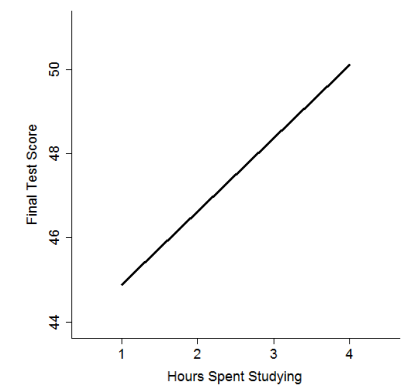
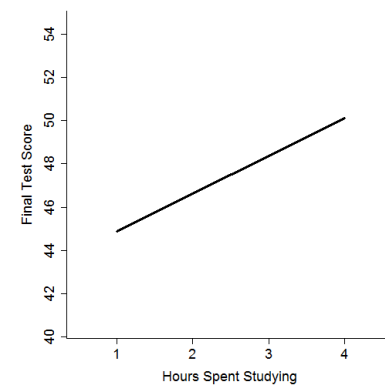
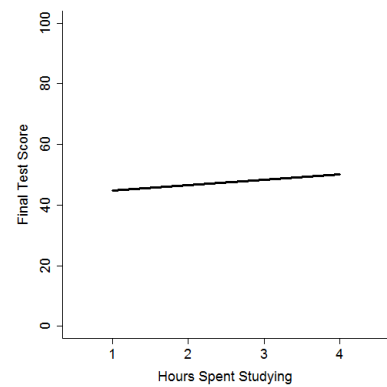
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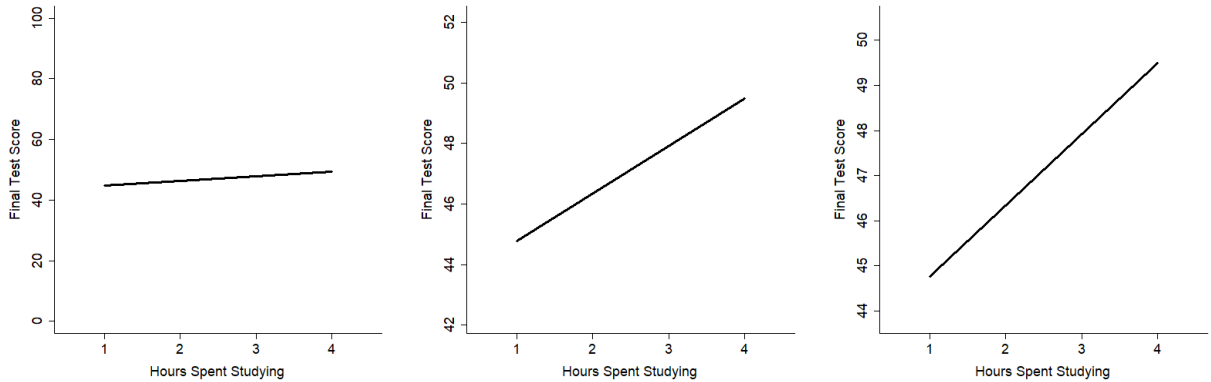


Figure A1. Sample stimuli for each of the 5 experiments. Each row corresponds to one experiment and shows a single set of a data plotted in the three different ways (full, standardized, and minimal). In all cases, the data show a medium effect (Cohen’s  $d = 0.5$ ). The number in parentheses under the experiment number indicates the range of the standardized condition.

For each graph, indicate if study style had 1. No effect, 2. A small effect, 3. A medium effect, 4. A big effect on final performance. Ready? Press ENTER”.

A trial began with a fixation cross at the center of the screen for 500ms. The graph was then shown. Above the graph, text reminded participants of the four response options. The graph remained until participants made a response, at which point, the graph disappeared and a blank screen was shown for 500ms. Each block of trials consisted of the presentation of each of the 120 graphs (3 graph types x 4 depicted effect sizes x 10 repetitions). Order was randomized within block, and participants completed 4 blocks for a total of 480 trials.

**Results and Discussion**

One participant only completed 431 trials, but their data were still included. The depicted effect size was recoded on a scale from 1 to 4 to be consistent with the scale of the response. The smallest effect size ( $d = .1$ ) was coded as 1.5 to account for the idea that this effect is smaller than a small effect but bigger than no effect. In later experiments, these graphs were replaced with graphs for which there was no effect instead of  $d = .1$ .

For each participant for each of the 3 axis range conditions, the data were submitted to separate linear regressions with estimated effect size as the dependent factor and actual effect size (recoded on a

scale from 1-4 then centered by subtracting 2.5) as the independent factor. The regressions produced two coefficients for each participant for each axis range condition. The slope indicates sensitivity to the size of the effect. A slope of 1 indicates perfect sensitivity. A slope less than 1 indicates attenuated sensitivity. The intercept indicates any bias to see effects as smaller or larger than their true size. One participant had slopes that were identified as outliers in the full and minimal conditions because they were greater than 1.5 times the interquartile range for each condition. This participant was excluded from the analysis (despite being the best performer in the group) because their data were not typical of the rest of the sample. Another participant had a slope less than 1.5 times the interquartile range in the full condition, and was also excluded for not being typical of the rest of the sample.

The coefficients were analyzed using paired-samples t-tests to compare each graph condition to the others. Analyses were done in R (R Core Team, 2017). Bayes factors were calculated using the BayesFactor package in R with a medium prior (Morey, Rouder, & Jamil, 2014). A Bayes factor greater than 3 indicates moderate evidence, and a Bayes factor greater than 10 indicates substantial evidence for the alternative hypothesis over the null hypothesis. Conversely, a Bayes factor less than .33 and less than .10 indicates moderate and substantial evidence for the null hypothesis over the alternative hypothesis. Effect sizes were calculated using the

recommendations of Lakens (2013), and 95% confidence intervals (CIs) on the effect size were calculated using the `cohen.d.ci` function in the PSYCH package (Revelle, 2018).

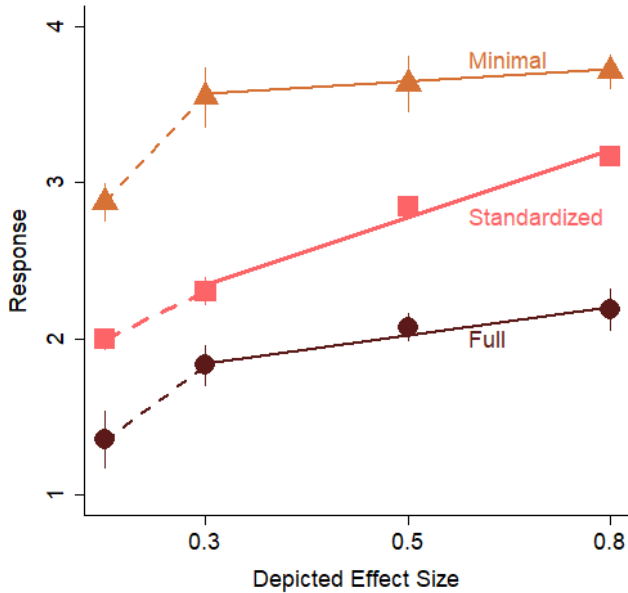


Figure A2. Mean response is plotted as a function of depicted effect size and graph type for Experiment 1. Error bars are 1 SEM calculated within-subjects. Solid lines represent linear regressions for depicted effects  $d \geq .3$ . Dashed lines represent linear regressions for depicted effects less than  $d \leq .3$ .

The standardized graphs produced significantly greater slopes than the full graphs,  $t(6) = 3.84$ ,  $p = .009$ ,  $d_z = 1.45$ , 95% CIs [.33, 2.51], Bayes factor = 7.54 (see Figure A2). With the standardized y-axis range, participants were more sensitive to the differences in actual effect size ( $M = .47$ ,  $SD = .11$ ) compared with graphs that showed the full range from 0 to 100 ( $M = .30$ ,  $SD = .07$ ). Sensitivity was also better for the standardized graphs than the minimal graphs,  $t(6) = 3.61$ ,  $p = .011$ ,  $d_z = 1.37$ , 95% CIs [.28, 2.40], Bayes factor = 6.17. The minimal graphs ( $M = .28$ ,  $SD = .04$ ) produced sensitivity similar to the full graphs,  $p = .51$ ,  $d_z = .26$ , 95% CIs [-.50, 1.01], Bayes factor = 0.43.

These data show an advantage for the standardized graphs because participants were more sensitive to differences among magnitudes of the depicted effect sizes with the standardized graphs than with the full or minimal graphs. However, the standardized graphs led to performance that was far

from perfect. The slope was .47, and perfect performance would have produced slopes of 1. Thus, even though the standardized graphs signify an improvement over the other two options, more work is still necessary to improve graph comprehension.

Another advantage for the standardized graphs can be seen with respect to bias. Bias scores were calculated as a percentage score of underestimation (negative values) and overestimation (positive values). They were calculated as the participant's coefficient for the intercept minus the true intercept (2.5) divided by the true intercept. There were significant differences between the bias scores across all conditions,  $ps < .003$ . The bias scores for the full graphs was negative ( $M = -27\%$ ,  $SD = 10\%$ ) and significantly below 0,  $t(6) = -7.01$ ,  $p < .001$ ,  $d_z = 2.64$ , 95% CIs [1.00, 4.27], Bayes factor = 82. The bias scores for the minimal graphs were positive ( $M = 36\%$ ,  $SD = 19\%$ ) and significantly above 0,  $t(6) = 4.91$ ,  $p = .003$ ,  $d_z = 1.86$ , 95% CIs [.57, 3.10], Bayes factor = 19. In contrast, the bias scores for the standardized graphs were significantly less biased than in the other conditions ( $ps < .003$ ), and were not significantly different from 0 ( $M = 1\%$ ,  $SD = 4\%$ ),  $t(6) = 0.47$ ,  $p = .66$ ,  $d_z = .18$ , 95% CIs [-.58, .82], Bayes Factor = 0.39. With the full graphs, most effects looked like small effects. Indeed, 91% of the trials with the full graphs were labeled as showing no effect or a small effect. With the minimal graphs, 58% of the effects were labeled as big effects and 88% were labeled as medium or big. With the standardized graphs, small effects looked small and medium effects looked medium (see Figure A3). However, the big effects only looked medium. Thus, the experiment was replicated but with a smaller range in the standardized condition to determine if that would improve detection of big effects.

## Experiment 2: Bar Graphs with Axis Range of 1.4 SD

Standardized graphs, for which the y-axis range is a function of the standard deviation, produced better sensitivity and less bias in participants who judged the size of the depicted effect compared with graphs that showed the full range and graphs that showed only the minimal range necessary to see the data. However, sensitivity with the standardized graphs was still below perfect performance. In this experiment, the range of the standardized graphs was decreased from 2 SDs to 1.4 SDs.

**Method**

Fourteen students in an introductory psychology course participated in exchange for course credit. Everything was the same in Experiment 1 except for the construction of the standardized graphs, for which the y-axis range went from the group mean minus 0.7 SD to the group mean plus 0.7 SD (see Figure A1). Thus, the standardized range was 1.4 SD (instead of 2 SD as in Experiment 1). In addition, the

simulated data were evaluated to ensure that the outcomes were similar to the intended outcomes. The effect size of the simulated data were compared to the intended effect size, and if they differed by more than 0.1, the data were resampled until the discrepancy was less than 0.1. Participants completed 4 blocks of 120 trials, and order was randomized within block.

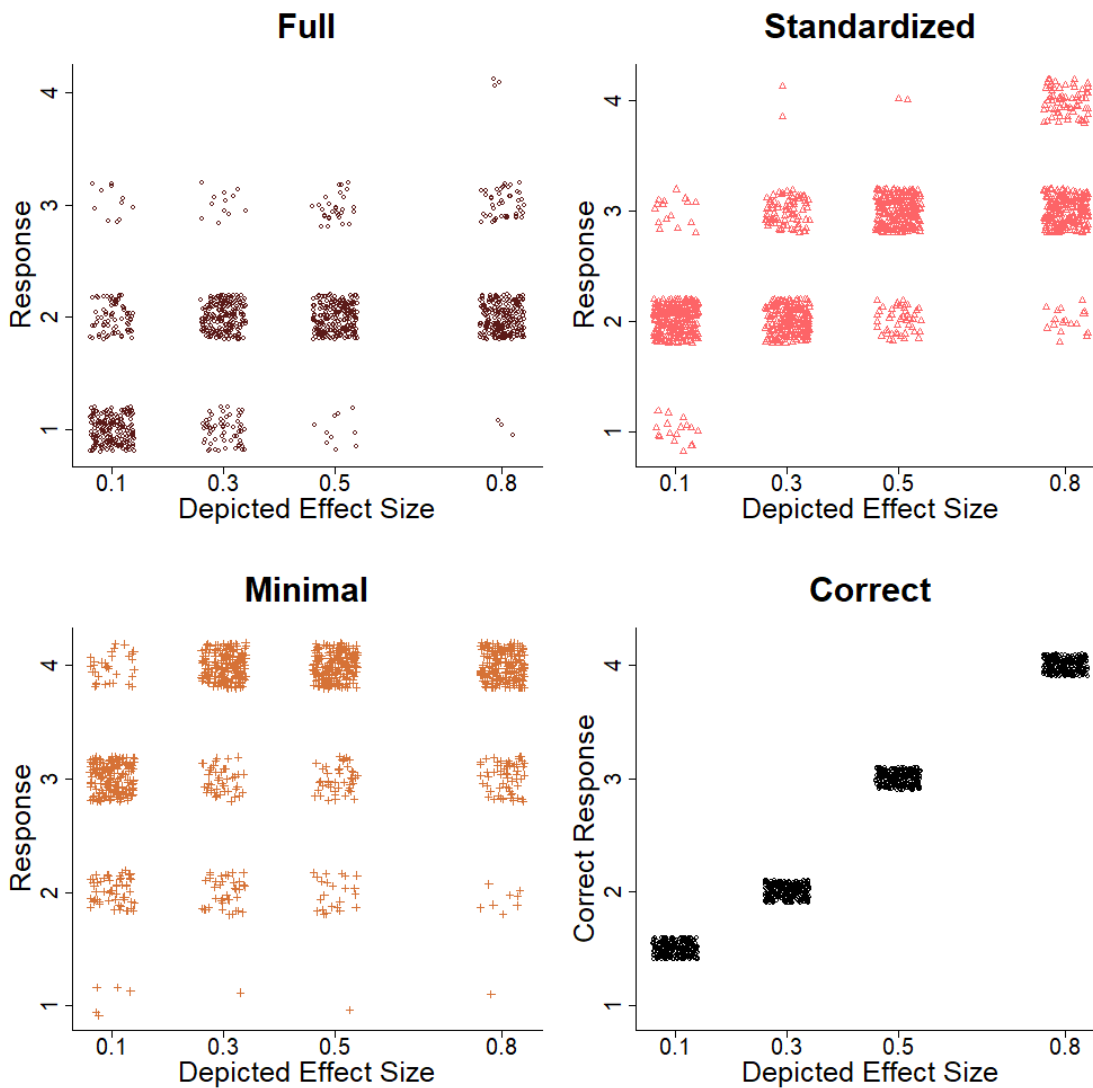


Figure A3. Response is plotted as a function of depicted effect size for the three types of axis range conditions (full, minimal, and standardized) for Experiment 1. The bottom right panel shows the correct response. Response was entered as 1 (no effect), 2 (small effect), 3 (medium effect), and 4 (big effect). Each point corresponds to one participant’s response on one trial. The data have been jittered along both axes to enable visibility.

## Results and Discussion

The data were analyzed as before. Three participants had a slope that was deemed an outlier for being beyond at least 1.5 times the interquartile range for the full or minimal graphs.

The slope, which indicates sensitivity to the size of the effect in the graph, was greater for the standardized graphs ( $M = .54$ ,  $SD = .17$ ) than the full graphs ( $M = .31$ ,  $SD = .08$ ),  $t(10) = 3.46$ ,  $p = .006$ ,  $d_z = 1.04$ , 95% CIs [.28, 1.77], Bayes factor = 9.00 (see Figure A4). Sensitivity was also greater for the standardized graphs than the minimal graphs ( $M = .30$ ,  $SD = .06$ ),  $t(10) = 4.07$ ,  $p = .002$ ,  $d_z = 1.23$ , 95% CIs [.42, 2.00], Bayes factor = 20. Replicating Experiment 1, the current data show that setting the range of the y-axis to be a function of the standard deviation, rather than the full range of options or the minimal range necessary to show the data, improved graph comprehension. Recall, participants were not asked to indicate how big the effect looked but rather how big the effect was. Full and minimal graphs both produced misleading impressions of the data that severely attenuated sensitivity to effect size. Simply setting the range of the y-axis in relation to the standard deviation improved readers' sensitivity to the data.

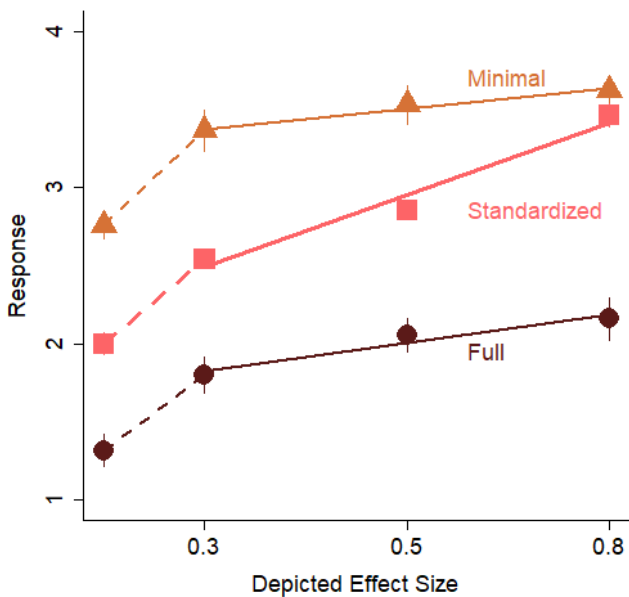


Figure A4. Mean response is plotted as a function of depicted effect size and graph type for Experiment 2. Error bars are 1 SEM calculated within-subjects. Solid lines represent linear regressions for depicted effects  $d \geq .3$ . Dashed lines represent linear regressions for depicted effects less than  $d \leq .3$ .

Bias was again found for the full and minimal graphs but not the standardized graphs. For the full graphs, the bias was to underestimate effect size by 28% ( $SD = 9\%$ ),  $t(10) = -10.51$ ,  $p < .001$ ,  $d_z = 3.17$ , 95% CIs [1.67, 4.64], Bayes factor > 100. Indeed, of all the trials with the full graphs, the effect was labeled as small or no effect on 90% of responses. The bias was of a similar magnitude but in the opposite direction for the minimal graphs,  $t(10) = 4.91$ ,  $p < .001$ ,  $d_z = 1.48$ , 95% CIs [.59, 2.33], Bayes factor = 61. With the minimal graphs, participants overestimated the size of the effects by 31% ( $SD = 21\%$ ). Over half of all effects with the minimal graphs were labeled big (53%), and 81% were labeled as medium or big. In contrast, the bias was much smaller ( $M = 6\%$ ,  $SD = 9\%$ ) for the standardized graphs, and only marginally significantly different from 0,  $t(10) = 2.13$ ,  $p = .059$ ,  $d_z = .64$ , 95% CIs [-.02, 1.28], Bayes factor = 1.50. The bias with the standardized graphs was far less than the biases observed with the full and minimal graphs,  $p < .001$ .

The evidence thus far is clear: graphs with a y-axis range that is a function of the standard deviation produces better sensitivity and less bias in participants when they are tasked with judging the size of an effect, compared with graphs that present the full range and with graphs that present only the minimal range necessary to view all of the data.

### Experiment 3: Bar Graphs with Error Bars

The graphs in Experiments 1 and 2 did not contain error bars. As a result, the graphs did not contain enough information to know if an effect was null, small, medium, or big. This was a conscious decision given that introductory psychology students might not know how to interpret error bars. Yet, it is necessary to know if standardized graphs still produce an advantage even when there is enough information presented in the graphs to be able to accurately answer the question. In addition, the graphs



with the smallest effects in Experiments 1 and 2 had the awkward feature of being bigger than no effect but smaller than a “small” effect, so it was unclear whether the correct answer should be 1 or 2. This ambiguity was eliminated in the current experiment.

**Method**

Thirteen students in an introductory psychology course participated in exchange for course credit.

Graphs were constructed similarly as in Experiment 2 with the following exceptions. The four effect sizes that were modeled were Cohen’s  $d = 0, .3, .5, \text{ and } .8$ , which corresponds to no effect, a small effect, a medium effect, and a big effect, respectively. The data were simulated as coming from two independent groups of 100 participants. The mean used to model the data for the hypothetical group that used the spaced studying strategy was always 50 (as in 50% accuracy on a memory test). The mean used to model the data for the hypothetical group that used the massed studying strategy was 50 minus 0, 3, 5, or 8 depending on the effect size being modeled. Using these means and a SD of 10, data were sampled from a normal distribution and summarized for the graphs. Error bars were calculated as 95% confidence intervals. In addition to the instructions given in Experiments 1 and 2, participants were also told the following: “Important! An effect is statistically significant if  $p < .05$ . However, you can also assess statistical significance by looking at error bars. Error bars are lines that extend from the mean of each condition. The mean of each condition is shown by the top of the bar. If the error bar from one condition overlaps the mean from the other condition, the effect is NOT significant. If neither bar overlaps the mean of the other condition, then the effect is significant. The farther apart the error bars, the bigger the effect.” Note that this rule of thumb is overly simplified. There can be cases for which the error bars overlap but the effect is statistically significant at the  $p < .05$  level (Cumming & Finch, 2005), but this level of nuance was not presented to the participants.

For each set of simulated data, 3 graphs were constructed. For the full graphs, the y-axis range went from 0 to 100. For the standardized graphs, the y-axis range went from the grand mean minus 0.6 SD to the grand mean plus 0.6 SD. For the minimal graphs, the bottom of the y-axis range was the

smallest combination of the mean minus the lower confidence interval minus 0.1, and the top of the range was the biggest combination of the mean plus the upper confidence interval plus 0.1. Participants completed 4 blocks of 120 randomized trials.

**Results and Discussion**

The data were analyzed as before. One participant had a negative slope for the standardized graphs, and another participant had a high slope for the full graphs. Both were 1.5 times beyond the interquartile range and excluded from analyses.

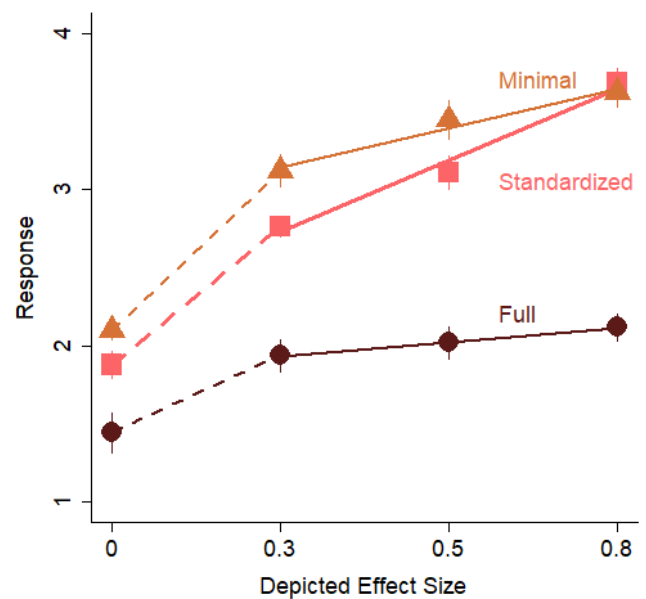


Figure A5. Mean response is plotted as a function of depicted effect size and graph type for Experiment 3. Error bars are 1 SEM calculated within-subjects. Solid lines represent linear regressions for depicted effects  $d \geq .3$ . Dashed lines represent linear regressions for depicted effects less than  $d \leq .3$ .

The slopes were steeper, showing better sensitivity, for the standardized graphs ( $M = .62, SD = .19$ ) compared with the full graphs ( $M = .24, SD = .09$ ) and the minimal graphs ( $M = .55, SD = .20$ ). The difference in slopes between the standardized and full graphs was significant,  $t(10) = 7.76, p < .001, dz = 2.34, 95\% \text{ CIs } [1.16, 3.50], \text{ Bayes factor } > 100$ . The difference in slopes between the standardized versus minimal graphs was also significant,  $t(10) = 3.09, p = .011, dz = .93, 95\% \text{ CIs } [.20, 1.63], \text{ Bayes factor } = 5.46$ . Even though all the information was the same across

the three graph conditions and even though this information was sufficient for determining the size of each effect, participants were better able to determine effect size when the range of the y-axis was a function of the standard deviation (see Figure A5).

The impression given by Figure A3 indicates that sensitivity was just as good if not better for the minimal graphs than the standardized graphs when comparing no effect to a small effect ( $d_s = 0$  and  $.3$ ), but sensitivity was better (steeper) for the standardized graphs when comparing across small, medium, and big effects ( $d_s = .3$ ,  $.5$ , and  $.8$ ). This impression prompted an unplanned analysis. Linear regressions were again conducted for each participant for each graph condition. However, in one set of regressions, only effect sizes  $0$  and  $.3$  were included. In another set of regressions, only effect sizes  $.3$ ,  $.5$ , and  $.8$  were included. Two additional participants were identified as outliers because the slopes for all three graphs in the latter analysis were 1.5 times beyond the interquartile range, and were excluded from the remaining analyses.

With respect to determining whether or not an effect is present (by comparing slopes for graphs depicting  $d_s = 0$  and  $.3$ ), all three graph types led to similar performance (Standardized:  $M = .89$ ,  $SD = .40$ ; Full:  $M = .49$ ,  $SD = .22$ ; Minimal:  $M = 1.02$ ,  $SD = .45$ ). With all three types of graphs, participants were sensitive to whether or not there was an effect, as shown by coefficients for each graph type that were positive and significantly greater than  $0$ ,  $p_s < .001$ . The standardized graph produced some benefit over the full graphs,  $t(8) = 2.82$ ,  $p = .022$ ,  $d_z = .85$ , 95% CIs  $[.14, 1.53]$ , Bayes factor = 3.76. The standardized graph was no better, and marginally worse, than the minimal graphs,  $t(8) = -1.82$ ,  $p = .11$ ,  $d_z = .55$ , 95% CIs  $[-.10, 1.17]$ , Bayes factor = 1.03. It should be noted that a bias to see all effects as being bigger (as found with minimal graphs) would lead to a steeper slope when comparing just the graphs that depict a null effect and a small effect. Thus, it cannot be known whether sensitivity is better with the minimal graphs or if the bias caused by the minimal graphs leads to greater estimates of sensitivity.

With respect to determining the magnitude of an effect that is present (by comparing slopes for graphs depicting  $d_s = .3$ ,  $.5$ , and  $.8$ ), the standardized graphs produced better sensitivity than the full or

minimal graphs (Standardized:  $M = .46$ ,  $SD = .11$ ; Full:  $M = .09$ ,  $SD = .06$ ; Minimal:  $M = .25$ ,  $SD = .13$ ),  $p_s \leq .001$ . The comparison between the standardized graphs to the full graphs resulted in a Bayes factor greater than 100,  $d_z = 2.81$ , 95% CIs  $[1.45, 4.14]$ . The comparison between the standardized graphs to the minimal graphs resulted in a Bayes factor of 65,  $d_z = 1.50$ , 95% CIs  $[.60, 2.35]$ . In each of the three graph types, participants showed some level of sensitivity to the magnitude of the effect, as shown by the coefficients being significantly greater than  $0$ ,  $p_s < .003$ .

In addition to better sensitivity with the standardized graphs, the standardized graphs also produced less bias compared with the other graphs,  $p_s \leq .001$ . For the full graphs, there was a 28% bias ( $SD = 12\%$ ) to underestimate effect size, which was significantly different from  $0$ ,  $t(10) = -7.82$ ,  $p < .001$ ,  $d_z = 2.36$ , 95% CIs  $[1.17, 3.52]$ , Bayes factor  $> 100$ . For the minimal graphs, there was a 14% bias ( $SD = 24\%$ ) to overestimate the size of the effect, which was marginally significantly from  $0$ ,  $t(10) = 2.03$ ,  $p = .069$ ,  $d_z = .61$ , 95% CIs  $[-.05, 1.25]$ , Bayes factor = 1.33. For the standardized effects, the bias was 7% ( $SD = 19\%$ ) and was not significantly different from  $0$ ,  $t(10) = 1.29$ ,  $p = .227$ ,  $d_z = .39$ , 95% CIs  $[-.23, .99]$ , Bayes factor = .58.

In summary, even with error bars, graphs with the y-axis range set as a function of the standard deviation produced better sensitivity and less bias compared with graphs that showed the full range and graphs that showed only the minimal range necessary to see the data.

#### **Experiment 4: Line Graphs with Axis Range of 1.4 SD**

The current experiment used line graphs as stimuli instead of bar graphs to see if the previous recommendations generalized to a different kind of graph.

#### **Method**

Twenty students in an introductory psychology course participated in exchange for course credit. Stimuli were graphs that were constructed by simulating data from two groups, and connecting their means with a line to create an impression of data across four groups. The four effect sizes that were



modeled were Cohen’s  $d = 0, .3, .5,$  and  $.8,$  which corresponds to no effect, a small effect, a medium effect, and a big effect, respectively. The y-axis range was full (0-100), minimal (smallest value minus 1 to largest value plus 1), or standardized (group mean minus 0.7 SD to the group mean plus 0.7 SD). Everything else was the same as in the previous experiments, except the x-axis was labeled as hours spent studying on a range from 1-4.

**Results and Discussion**

The data are shown in Figure A6. The data were analyzed as before with three separate linear regressions for each participant for each graph type for each combination of all effect sizes,  $d = 0$  and  $.3$  only, and  $d = .3 - .8$  only. One participant had slopes greater than 1.5 times the interquartile range for the full and minimal graphs, and 3 participants had slopes less than 1.5 times the interquartile range for the minimal graphs. All 4 were excluded.

For regressions on all effect sizes depicted in the graphs, the standardized graphs lead to greater slopes than the full graphs,  $t(15) = 7.16, p < .001, dz = 1.79,$  95% CIs [.98, 2.59], Bayes factor  $> 100$  (see Table A1). The standardized graphs did not lead to significantly different slopes than the minimal graphs when calculated across the entire range,  $t(15) = 0.18, p = .86, dz = .05,$  95% CIs [-.45, .53], Bayes factor = .26. However, this is because the minimal graphs produced superior performance with respect to determining whether there was an effect or not but inferior performance when an effect was present and the magnitude had to be determined. For regressions comparing  $d = 0$  to  $d = .3,$  the slopes for the minimal graphs were higher than for the standardized graphs,  $t(15) = -4.70, p < .001, dz = 1.17,$  95% CIs [.52, 1.81], Bayes factor  $> 100.$  Again, recall that the bias generated by the minimal graphs to see effects as bigger would produce greater sensitivity scores even if participants were not necessarily more sensitive to the effect. Indeed, the slope coefficient is 1.29, which is greater than perfect accuracy of 1, which implies some bias. For regressions comparing  $ds > 0,$  the slopes for the standardized graphs were higher than for the minimal graphs,  $t(15) = 3.05, p = .008, dz = .76,$  95% CIs [.19, 1.31], Bayes factor = 6.46. This suggests that the standardized graphs still produced better outcomes than the full or minimal graphs.

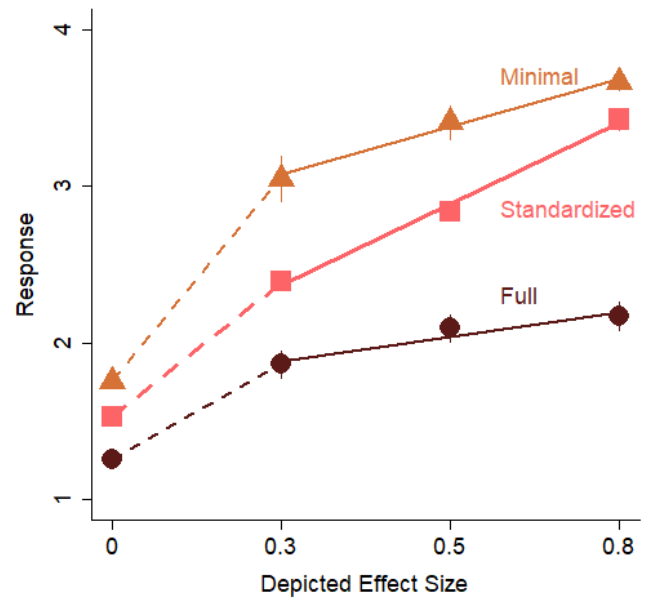


Figure A6. Mean response is plotted as a function of depicted effect size and graph type for Experiment 4. Error bars are 1 SEM calculated within-subjects. Solid lines represent linear regressions for depicted effects  $d \geq .3.$  Dashed lines represent linear regressions for depicted effects less than  $d \leq .3.$

Table A1.

Mean (and SD) coefficients for the slopes for each graph type for each analysis from Experiment 4.

Graph Type	All data	$ds = .3-.8$	$ds = 0 - .3$
Full	.30 (.08)	.15 (.08)	.61 (.22)
Standardized	.61 (.15)	.52 (.18)	.86 (.32)
Minimal	.61 (.06)	.31 (.25)	1.29 (.56)

Note. The slopes indicate the linear relationship between the size of the effect depicted and the estimate of the effect size, both of which were coded on a scale from 1-4.

Regarding bias, similar results were found as in previous experiments. The bias was -26% (SD = 11%) with the full graphs, indicating a bias to underestimate the effects,  $t(15) = -9.52, p < .001, dz = 2.38,$  95% CIs [1.39, 3.34], Bayes factor  $> 100.$  The bias was 19% (SD = 17%) with the minimal graphs, indicating a bias

to overestimate the size of the effects,  $t(15) = 4.36$ ,  $p < .001$ ,  $d_z = 1.09$ , 95% CIs [.46, 1.70], Bayes factor = 64. With the standardized graphs, the bias was 2% (SD = 10%), which was not significantly different from 0,  $t(14) = 0.73$ ,  $p = .48$ ,  $d_z = .18$ , 95% CIs [-.31, .67], Bayes factor = .32. With the line graphs, as with the bar graphs, the standardized axis range produced better sensitivity and less bias than the full axis range or the minimal axis range.

### Experiment 5: Line Graphs with Axis Range of 1 SD

The current experiment replicated Experiment 4 using a smaller axis range for the standardized graphs.

#### Method

Fifteen students in an introductory psychology course participated in exchange for course credit. The stimuli were the same as in Experiment 4 except that for the standardized graphs, the range was the group mean  $\pm 0.5$  SD.

#### Results and Discussion

The data were analyzed as before with three separate linear regressions for each participant for each graph type for each combination of all effect sizes,  $d = 0$  and  $.3$  only, and  $d = .3 - .8$  only. One participant had a slope that was less than 1.5 times the interquartile range for the minimal graphs, and one had a slope greater than 1.5 times the interquartile range for the full graphs. Both were excluded. The mean slope coefficients for the remaining participants are shown in Table A2 and the data are shown in Figure A7.

Table A2.

Mean (and SD) coefficients for the slopes for each graph type for each analysis from Experiment 5.

Graph Type	All data	$d_s = .3-.8$	$d_s = 0 - .3$
Full	.32 (.12)	.20 (.13)	.58 (.22)
Standardized	.55 (.20)	.48 (.17)	.76 (.49)
Minimal	.53 (.16)	.36 (.25)	.96 (.52)

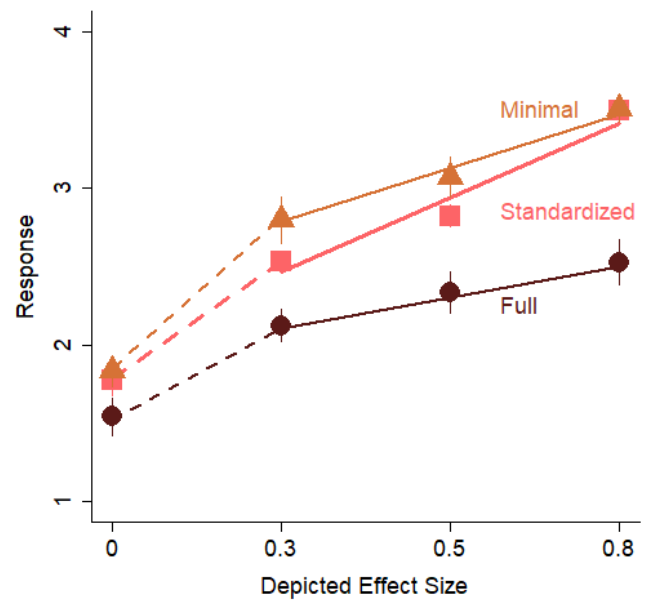


Figure A7. Mean response is plotted as a function of depicted effect size and graph type for Experiment 5. Error bars are 1 SEM calculated within-subjects. Solid lines represent linear regressions for depicted effects  $d \geq .3$ . Dashed lines represent linear regressions for depicted effects less than  $d \leq .3$ .

The patterns match those found in Experiment 4. Participants were more sensitive to the size of the effect for the standardized graphs than for the full graphs when all trials were included,  $t(13) = 4.41$ ,  $p < .001$ ,  $d_z = 1.22$ , 95% CIs [.48, 1.94], Bayes factor = 46, and when trials for which the effect size depicted was greater than 0,  $t(13) = 6.69$ ,  $p < .001$ ,  $d_z = 1.86$ , 95% CIs [.93, 2.76], Bayes factor > 100, but not when only trials for which the effect size depicted was null or small,  $t(13) = 1.41$ ,  $p = .19$ ,  $d_z = .39$ , 95% CIs [-.18, .95], Bayes factor = .62. Participants were more sensitive to the size of the effect for the standardized graphs than for the minimal graphs but only when the depicted effect in the graph was greater than 0,  $t(13) = 2.59$ ,  $p = .023$ ,  $d_z = .72$ , 95% CIs [.09, 1.32], Bayes factor = 2.88. There was no difference in sensitivity across all effect sizes,  $p = .65$ , Bayes factor = .31, and the minimal graphs produced better sensitivity when only data from graphs depicting a null or small effect were included,  $t(13) = -3.11$ ,  $p = .009$ ,  $d_z = .86$ , 95% CIs [.21, 1.49], Bayes factor = 6.27. As before, the bias created by the minimal graphs could account for this apparent increase in sensitivity.

Regarding the bias, the full graphs produced a bias of -15% (SD = 17%), indicating a bias to underestimate effect size,  $t(12) = -3.07$ ,  $p = .010$ ,  $d_z = .85$ , 95% CIs [.20, 1.48], Bayes factor = 5.90. The minimal graphs produced a bias of 12%, (SD = 20%), which was marginally above 0,  $t(12) = 2.21$ ,  $p = .047$ ,  $d_z = .61$ , 95% CIs [.01, 1.20], Bayes factor = 1.69. The standardized graphs led to a small bias of 6% (SD = 14%), that was marginally close to 0,  $t(12) = 1.63$ ,  $p = .13$ ,  $d_z = .45$ , 95% CIs [-.13, 1.01], Bayes factor = .80.

**Across Experiment Comparisons**

Sample size was not selected to achieve sufficient power to do analyses across experiments. To facilitate preliminary exploration of the data, the coefficients are reported in Tables S3, S4, and S5, and are plotted in Figure A8 and Figures 2 and 3 in the main Table A3.

text. It may be interesting to note that sensitivity of the size of the effect was not notably better with error bars than without error bars even though error bars are necessary to understand effect size. Although this may not be surprising given the participants being introductory psychology student, the pattern is consistent with previous findings that many researchers do not know how to interpret error bars (Belia, Fidler, Williams, & Cumming, 2005). In addition, the lack of noticeable differences in sensitivity between the experiments suggests that the use of a y-axis range that is approximately 1.5 SDs could help better report the results in cases for which researchers neglect to include error bars.

*Mean slopes (and standard deviations) from regressions on all trials for each of the 5 experiments.*

Graph Type	Exp 1	Exp 2	Exp 3	Exp 4	Exp 5
Full	.28 (.10)	.31 (.08)	.21 (.08)	.30 (.08)	.32 (.13)
Standardized	.46 (.13)	.54 (.17)	.58 (.17)	.61 (.15)	.55 (.20)
Minimal	.27 (.03)	.30 (.06)	.49 (.15)	.61 (.06)	.53 (.16)

Note. A slope of 1 indicates perfect performance and a slope of 0 indicates chance performance.

Table A4.

*Mean slopes (and standard deviations) from regressions on trials for which Cohen's d > 0.1 for each of the 5 experiments.*

Graph Type	Exp 1	Exp 2	Exp 3	Exp 4	Exp 5
Full	.17 (.10)	.18 (.10)	.09 (.06)	.15 (.08)	.20 (.13)
Standardized	.42 (.16)	.46 (.16)	.46 (.11)	.52 (.18)	.48 (.17)
Minimal	.07 (.09)	.13 (.14)	.25 (.13)	.31 (.25)	.36 (.25)

Note. A slope of 1 indicates perfect performance and a slope of 0 indicates chance performance.

Table A5.

*Mean bias scores as a percentage (and standard deviations) for each of the 5 experiments.*

Graph Type	Exp 1	Exp 2	Exp 3	Exp 4	Exp 5
Full	-27 (5)	-28 (9)	-25 (10)	-26 (11)	-15 (17)
Standardized 1 (4)		6 (9)	14 (13)	2 (10)	6 (14)
Minimal	36 (21)	31 (21)	23 (16)	19 (17)	12 (20)

Note. Bias scores were calculated as a percent bias based on intercepts from regressions on all trials including those for which Cohen's  $d = 0$ .