

No Myth far and wide: Relay Swimming is Faster than Individual Swimming and the Conclusion of Skorski et al. (2016) is Unfounded

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Skorski, Extebarria, and Thompson (2016) aim at our article on relay swimmers (Hüffmeier, Krumm, Kanthak, & Hertel, 2012). We have shown that professional freestyle swimmers at relay positions 2 to 4 swam faster in the relay than in the individual competition if they had a high chance to win a relay medal. After applying a reaction-time correction that controls for different starting procedures in relay and individual competitions, Skorski et al. (2016) conclude that swimmers in relays do not swim faster. At first sight, their results appear to show this very pattern. However, we argue that the authors' findings and conclusion—that our finding is a myth—are not warranted. First, we have also controlled for quicker reaction times in the relay competition. Our correction has been based on the swimmers' own reaction time data rather than on a constant reaction time estimate and is, thus, more precise than theirs. Second, Skorski et al. treat data from international and national competitions equally although national relay competitions are less attractive for the swimmers than national individual competitions. This difference likely biases their data towards slower relay times. Third, the authors select a small and arbitrary sample without explicit power considerations or a clear stopping rule. Fourth, they unfavorably aggregate their data. We conclude that the reported results are most likely due to the methodological choices by Skorski et al. and do not invalidate our findings.

Keywords: teams; groups; performance gains; effort gains; relay swimming; reaction-time correction; exchange-block times

In their recent study, Skorski, Extebarria, and Thompson (2016) set out to show that our findings (Hüffmeier, Krumm, Kanthak, & Hertel, 2012) are unfounded or, in their words, “a myth” (cf. the title of their article). In the current commentary, we intend to illustrate (i) that the reported findings in the article by Skorski et al. (2016) most likely result from unwarranted methodological choices and (ii) that their assessment of our findings is, thus, not valid. To do so, we first portray our pertinent studies. We focus on our study procedures, the methodological choices we made, our main findings, our explanation as well as possible alternative explanations for these findings, and we also discuss the robustness of our findings in relation to different methodological choices. We then describe the procedure and findings of Skorski et al. (2016). Finally, we point out four central shortcomings in the Skorski et al. article and conclude that the results reported by Skorski et al. cannot effectively question our results.

Our Studies

Main goal of our studies.

Our article (Hüffmeier et al., 2012) is part of a series of published (Hüffmeier et al., 2017; Hüffmeier & Hertel, 2011; Hüffmeier, Kanthak, & Hertel, 2013) and currently not yet published studies (Schleu, Mojzisch, & Hüffmeier, 2018). They build on a long research tradition in (social) psychology that investigated—mostly in laboratory experiments—how working as part of a group affects the members’ motivation and performance (see Karau & Williams, 1993; Weber & Hertel, 2007, for overviews). In their meta-analysis, Weber and Hertel (2007) have shown that group work can result in *effort gains in groups* (i.e., higher effort during group as compared to individual work).¹ They have further found that perceived social indispensability (i.e., the perception that the own contribution to the group performance is critical) is an important driver of such effort gains. Our studies investigated whether and under which conditions group work results in such indispensability-based effort gains in groups *outside the research*

¹ Prior research used the term “motivation gains in groups”. However, motivation encompasses the direction, intensity, and persistence of behavior. Pertinent small group research such as ours typically does not study the direction, but the

laboratory (see Hüffmeier & Hertel, *in press*, for an overview of the swimming studies).

Methodological approach of our studies.

Most of our pertinent studies were situated in a professional swimming context and compared swimming performances from the individual and relay competitions of the same event (e.g., the 2008 Olympic Games; in a recent study, we investigated performances in a track and field context, cf. Schleu et al., 2018). Due to the high standardization of performances, professional swimming is an apt context to study the motivational impact of group versus individual work. In fact, only the prescribed starting procedures differentiate between the work that swimmers do in the individual and relay competitions.

The starting rules for official swimming competitions (cf. the rules of the world swimming federation [FINA]) prescribe the respective starting procedures:² Swimmers at the first relay position and swimmers in the individual competition have to follow identical starting rules. These swimmers perform a “flat start”, meaning that they cannot reliably anticipate their acoustic starting signal and are not allowed to move before this starting signal sounds. By contrast, in the relay competition, swimmers at relay positions 2 to 4 perform a “flying start”. That is, they can reliably anticipate their “starting signal” (i.e., their predecessor touching the pool wall). And they are allowed to move before this starting signal as long as at least one foot still touches the starting block when their predecessor touches the wall. Obviously, these differences lead to quicker starting or reaction times in the relay as compared to the individual competitions for swimmers at relay positions 2 to 4. Consequently, the resulting reaction-time advantages in relays of about 0.422 to 0.524 sec (depending on the swimmers’ gender, relay position, and swimming distance, see Hüffmeier et al., 2017; see also Table 1) have to be controlled to assess the pure swimming time difference between relay and individual swimming.

Our studies shared the following approach: First, we sampled data of swimmers taking part in both the individual and relay competitions in the same swimming

intensity and persistence of behavior (i.e., the two components of effort). We, thus, prefer the term “effort gains in groups”.

² See the FINA homepage:

<http://www.fina.org/content/fina-rules>

Table 1. Overview of the Reaction Time Estimates Applied in Skorski et al. (2016) and in the two Studies of Hüffmeier et al. (2017) for the Four Relay Positions (Reported in Seconds)

| Relay position | Skorski, Extebarria, & Thompson (2016) ¹ | Hüffmeier et al. (2017), Study 1 ² | Hüffmeier et al. (2017) Study 2 ³ |
|----------------|---|---|--|
| Position 1 | -- | 0.019 s (95% CI: -0.07; 0.045) | -0.006 s (95% CI: -0.010; -0.002) |
| Position 2 | 0.48 s | 0.505 s (95% CI: 0.442; 0.568) | 0.465 s (95% CI: 0.448; 0.482) |
| Position 3 | 0.48 s | 0.455 s (95% CI: 0.366; 0.544) | 0.455 s (95% CI: 0.431; 0.479) |
| Position 4 | 0.48 s | 0.524 s (95% CI: 0.445; 0.603) | 0.463 s (95% CI: 0.447; 0.479) |

Note. ¹ Skorski et al. (2016) applied a constant of 0.48 seconds to the swimming times of 100 meter relay swimmers at the second, third, and fourth relay position as a reaction time estimate. This constant was taken from a swimmer sample that was unrelated to their own sample (see Saavedra et al., 2014, for details).

² In their Study 1, Hüffmeier et al. (2017) collected data to specifically estimate the reaction time advantages in relays separately for relay position, swimming distance, and gender. To illustrate the methodological differences as compared to Skorski et al. (2016), the data from men competing in the 100 meter freestyle competition is reported in Table 1.

³ In their Study 2, Hüffmeier et al. (2017) used a data set not specifically collected for this purpose to derive reaction time estimates. To illustrate the methodological differences as compared to Skorski et al. (2016), the data from men competing in the 100 meter freestyle competition is reported in Table 1.

discipline at the same professional and international events (e.g., 100 meter freestyle competitions at the 2008 Olympic Games).³ We analyzed data from national competitions in only one article (Hüffmeier et al., 2017). Second, with one exception (Hüffmeier & Hertel, 2011), we made sure that the included data came from comparable stages in both competitions (e.g., final races from both the individual and relay competitions). Third, in most studies, we only included a swimmer if his/her reaction times from both competitions were available (Hüffmeier et al., 2013, 2012; Hüffmeier & Hertel, 2011). As an alternative approach, we conducted a pilot study to derive estimates of the reaction-time advantage separately for each relay position (Hüffmeier et al., 2017; see also Table 1). We used both types of data (i.e., original reaction-time data or estimates) to correct for the reaction-time advantages separately for each relay position (i.e., by subtracting the respective

reaction times from the swimming times for the individual and relay competition).

Fourth, we either included all data that was available at the time of a study (e.g., Study 1 in Hüffmeier et al., 2017; Hüffmeier et al., 2013; 2012) or we conducted a power analysis to determine the necessary sample size to detect the expected effects (Study 2 in Hüffmeier et al., 2017). Fifth, in accordance with our finding that indispensability perceptions increase across the positions of the relay (Hüffmeier & Hertel, 2011), we expected a parallel increase in effort across the relay. Thus, we conducted analyses testing a corresponding linear contrast as well as separate analyses for each relay position rather than aggregating the swimming performance data across relay positions. Sixth, we interpreted the observed performance differences between the individual and relay competition as differences in the swimmers' effort levels.

³ The resulting data are not distributed equally across relay positions because normally only two swimmers from one country can qualify for individual competitions at the prestigious international championships and these two

strongest national swimmers are typically assigned to the first and fourth relay position (see, for instance, Table 2 in Neugart & Richiardi, 2013, or Table 1 below).

Results of our studies.

Across all five studies focusing on swimming performances, we found the expected linear increase in effort across the relay—in those conditions where we expected it. The effect sizes for this linear increase were typically medium to large (Cohen, 1992; see also Table 2). In four of five studies (except for Study 1 in Hüffmeier et al., 2017), the effort level at the first relay position was comparable to the expended effort in the individual competition (i.e., no performance differences between the two competition types). Again, in four of five studies (except for Study 2 in Hüffmeier et al., 2017), we observed very small, but statistically significant effort gains in groups at relay position 2 (i.e., quicker swimming times in the relay as compared to the individual competition; see Table 2). We consistently found effort gains at position 3, which were largely comparable in size to those at position 2. At the fourth position, we again consistently found effort gains. These gains were more pronounced than at the prior positions

in all but one study (Hüffmeier et al., 2013; see Table 2), but they were still small in size.

Concerning potential moderators of these findings, we have initial evidence that type of swimming relay (e.g., freestyle vs. medley relays) may act as a moderator (i.e., we found no effort gains in medley relays, Hüffmeier et al., 2013). We have more reliable evidence that (i) the swimmers' medal chances in the relay and (ii) the relative valence of the obtainable group outcomes may moderate these findings in the following way: Swimmers without chances to win a medal in the relay competition and swimmers competing for less attractive outcomes in the relay than in the individual competition did not exhibit effort gains in groups (Hüffmeier et al., 2017; 2012).

By contrast, we did not find that swimmers' gender, the swimming distance (100 vs. 200 meters) or the sports discipline (swimming vs. track and field relays) moderated the observed findings. In a first track and field study, we studied runners' efforts ($N = 397$) in the individual and relay 400 meter running competitions of

Table 2. Summary of Exemplary Previous Studies and Effect Sizes Indicating Increases in Effort Across the Relay and Effort Gains in Groups

| Study | Total Sample Size | Relevant Condition | Sub-sample Size | Effect Sizes ¹ | | | | Planned contrast ⁴ |
|-----------------------------------|-------------------|---|-----------------|---------------------------|------------|------------|------------|-------------------------------|
| | | | | Position 1 | Position 2 | Position 3 | Position 4 | |
| Hüffmeier et al. (2012) | 199 | High instrumentality of group performance ² | 151 | -0.014 | 0.040 | 0.045 | 0.107 | .74 |
| Hüffmeier et al. (2017) – Study 1 | 302,576 | High instrumentality of group performance ² and high valence competitions ³ | 928 | 0.010 | 0.003 | 0.007 | 0.089 | .80 |

Note. To allow for a high comparability of the findings in the 2017 data set with the original findings in Hüffmeier et al. (2012), only 100 m free-style races were analyzed for this table. Effect sizes are Cohen's d , using the standard deviation of individual swimming times. Negative effect sizes indicate performance losses.

¹ Mean difference between individual and relay competition times (corrected by reaction times of different starting procedures) divided by the standard deviation of the individual times.

² Good chance of winning a medal.

³ Olympic Games, World Championships, European Championships, Pan Pacific Games, Commonwealth Games, and Universiades.

⁴ Contrary to the analyses in the referenced papers, planned contrasts were analyzed within the relevant condition only. The t -value of the planned contrast was converted into Cohen's d .

the same prestigious and international sports events (Schleu et al., 2018). Paralleling our findings from the swimming domain, we found a linear increase in self-reported effort and also in objective performance across the running relay. We again observed effort gains in groups for the athletes running at later positions in the relay.

Explanation for our findings.

To learn whether the perceived indispensability of the own contribution to the group performance (cf. Weber & Hertel, 2007) could also be a relevant mediator explaining the effort gains of relay swimmers, we conducted a study with competitive adolescent swimmers (Hüffmeier & Hertel, 2011). Confirming our expectations, they perceived an increasing indispensability of the own contribution to the group performance across the relay (i.e., they progressively perceived that an own bad performance could not be compensated by fellow relay swimmers). Because this increase in indispensability across the relay parallels the consistently observed increase in effort across the relay, we believe that perceived indispensability may be an important mechanism underlying the effort gains among relay swimmers.

Alternative explanations.

In this section, we discuss four prominent alternative explanations for our main findings. First, the different starting procedures in the individual and relay condition most likely do not represent a valid alternative explanation: This account cannot explain the consistently found linear increase in effort across the relay and, in particular, the most pronounced effort gains at the last relay position. Second, the order in which swimmers took part in their competitions (e.g., relay competition first vs. individual competition first) could be proposed as another explanation. In the studies where we could test this explanation, we did, however, not observe a significant moderation by competition order (Hüffmeier et al., 2017; 2012).

Third, although prominent in the literature on effort gains in groups (Weber & Hertel, 2007), social comparison is a rather unlikely alternative explanation. Obviously, the clearest possibilities for comparing one's performance with those of others exist in the individual competition and for the first relay swimmers because all swimmers start simultaneously in these situations. However, first relay swimmers did not exhibit effort gains in any of our studies and the backlogs or leads accumulated over the course of the relay competitions

impede such comparisons for the remaining relay positions. These are, however, the positions where effort gains were observed.

Fourth, group members' relative strength was discussed as an alternative explanation (cf. Osborn, Irwin, Skogsberg, & Feltz, 2012), meaning that effort gains were primarily expected among the weaker swimmers of a relay. This expectation is transferred from laboratory research where the weaker group members often exhibit effort gains in typical laboratory tasks (cf. Weber & Hertel, 2007). We conducted a study (Study 2 in Hüffmeier et al., 2017) to put this explanation to an empirical test. To do so, we collected data on the relative strength of all relay members and ordered them accordingly in our analysis. The resulting linear contrast for relative member strength was, however, not significant and could therefore not explain the observed results.

Robustness of our findings in relation to different methodological choices.

Our extant results were largely unaffected by the type of analysis and the type of reaction-time correction we used. In most of our studies, we report analyses using difference scores (i.e., we subtracted the relay from the individual swimming times). However, our results did not hinge on this type of analysis: We obtained equivalent results in all studies when conducting repeated-measures analyses.

Our results were also unaffected by the different types of data we applied to correct for the reaction-time advantages in the relay. First, we observed parallel findings when using the exact reaction-time data of the swimmers (e.g., Hüffmeier & Hertel, 2011; Hüffmeier et al., 2012) and when using reaction-time estimates (Hüffmeier et al., 2017). Second, different types of reaction-time estimates did also not affect our results: In our two most recent studies, we applied two sets of reaction-time corrections (i.e., from a pilot study designed to establish estimates and from a big data set not specifically collected for this purpose; cf. Hüffmeier et al., 2017, and Table 1). We obtained comparable results with both sets of reaction time corrections.

We are, however, well aware that our findings—especially at relay positions 2 and 3—are probably sensitive to methodological choices because the related effects are very small (see Table 2). For instance, if a uniform and higher estimate is applied to each relay position to correct for the reaction-time advantages in relays (as, for instance, in Neugart & Richiardi, 2013) we probably would not observe effort gains at these positions. Applying a uniform estimate to each relay position and choosing higher estimates are, however,

doubtful methodological decisions (see Hüffmeier et al., 2017, for related data and a discussion; see also Table 1). Nevertheless, the effort gains at relay positions 2 and 3 are probably less robust than the effort gains at the fourth relay position and especially less robust than the overall linear increase in effort across the relay.

Our study targeted by Skorski et al. (2016).

In our study (Hüffmeier et al., 2012), we analyzed the swimming performances of $N = 199$ professional freestyle swimmers from about a decade of 100 meter competitions (i.e., Olympic Games from 1996 to 2008, World Championships from 1998 to 2011, and European Championships from 2000 to 2010) and we corrected their swimming times from the individual and relay competition by using their own reaction times. Our analysis shows that swimmers at the first relay position swam equally fast in both competition types while swimmers at relay positions 2 to 4 swam faster in the relay as compared to the individual competition if they had a high chance to win a relay medal.⁴ These effort gains in groups increased from the second to the fourth relay position.

The Skorski et al. (2016) Article

Skorski et al. (2016) analyze the swimming performances of 166 male swimmers from 778 races. These races were either part of the 100 and 200 meter freestyle competitions of the Olympic Games from 2000, 2004, and 2012 ($n = 144$ races) or of international and national Non-Olympic 100 and 200 meter competitions ($n = 634$ races). The authors frame their use of a reaction-time correction controlling for the quicker reactions times in the relay as compared to the individual competition as the main contribution of their research.

The authors report that they do no longer find that the analyzed swimmers swim faster in relay events than in their individual races if they control for the existing reaction-time advantages in the relay competitions. As corrections, they add 0.48 sec to the swimming times of the 100 m relay competitions and 0.45 sec to the swimming times of the 200 m relay competitions. While the authors' conclusion does not seem to be fully accurate (on page 412, the authors report that "there was a small positive performance effect after adjusting for reaction time" for the swimming times of swimmers at positions

2 to 4 in Olympic 100 meter freestyle relays), we appreciate their efforts to critically assess our own findings. Upon closer inspection, however, it seems that the authors may have overlooked some aspects of our study and that they made several methodological decisions that taint their reported results and question their conclusion.

Clarification and Four Central Shortcomings of the Skorski et al. (2016) Article

A clarification of the research literature: Most pertinent studies applied a reaction-time correction.

Skorski et al. present their use of a reaction-time correction as a significant contribution. This creates the impression to readers that we did not apply a reaction-time correction in the targeted article (Hüffmeier et al., 2012). But of course, we and others have consistently done so in the targeted and other pertinent articles (e.g., Hüffmeier & Hertel, 2011; Hüffmeier et al., 2017; 2013; Neugart & Richiardi, 2013; Osborn et al., 2012).

The first shortcoming: Our applied reaction-time correction is more precise than that of Skorski et al. (2016).

The reaction-time correction in our 2012 article was more precise than the one reported by Skorski et al. (2016) because we relied on the swimmers' own reaction time data: Thus, we did not use reaction-time estimates, but only included swimming data for a swimmer in our analysis if his/her exact reaction times were available for both the individual and relay competition. By contrast, Skorski et al. (2016) added two different and rather rough constants to their data from the 100 and 200 meter relay competitions (0.48 sec and 0.45 sec, respectively). These constants are not related to the swimmers of their sample, but were instead derived from a prior, unrelated study (cf. Saveedra, García-Hermoso, Escalante, Dominguez, Arellano, & Navarro, 2014).

Our own recent research suggests that adding the same constant to each relay swimming time is an imprecise approach because reaction times differ depending on the relay position (see Hüffmeier et al., 2017 and Table 1, for details). Thus, the authors did not only ignore that their use of a reaction-time correction was not

⁴ This finding is not restricted to relay swimming with high medal chances as compared to individual swimming with freely varying medal chances. We have obtained similar

findings when the chances of success in both the relay and individual competition varied freely (see Study 2 in Hüffmeier et al., 2017).

novel at all, but they also applied a less precise reaction-time correction than we did. It is, thus, unlikely that the absence of observed differences between swimming times from the relay and individual competitions is due to the mere use of reaction-time corrections as the authors claim. In contrast, their imprecise correction may possibly have obscured otherwise observable differences.

The second shortcoming: The sample in Skorski et al. (2016) with many more Non-Olympic than Olympic races systematically biases the results towards slower relay swimming times.

Skorski et al. (2016) analyze a variety of competitions: Races from Olympic competitions ($n = 144$) as well as from international and national Non-Olympic competitions ($n = 634$), including national championships. If the authors attempted to show that the results of our article are unfounded, this sampling approach is problematic because we focused our analysis exclusively on *international championships* (i.e., Olympic Games, World and European Championships; see Hüffmeier et al., 2012).

In these championships, the relay and individual competitions are equally attractive for the professional swimmers—in both types of competitions they pursue the ultimate goal to win the race or at least to win a medal. The situation is different for *national championships*: Here, the individual competition is much more attractive than the relay competition because swimmers can win and qualify for the upcoming international events in the individual competition (e.g., Olympic Games). By contrast, they can typically *only win, but not qualify* in the relay competition (i.e., relays at international events mostly consist of the best four national swimmers and not of the best relay from the national championship). In fact, many national swimming associations do not even consider relay performances from national championships as a relevant criterion for selecting swimmers for major international swimming competitions like the Olympic Games or World Cham-

pionships.⁵ We have found empirical evidence that accords with the differing attractiveness of national individual and relay competitions in a re-analysis of a data set with more than 300,000 races (Hüffmeier et al., 2017; cf. Neugart & Richiardi, 2013). In this sample, professional athletes show the typical pattern of increasing effort gains in relays from position 2 to 4 for international competitions, while we even observe effort losses in groups for national and local competitions (i.e., slower swimming times in the relay as compared to the individual competition; for a related hypothesis on the valence of group outcomes, see Karau, Markus, & Williams, 2000).

In Skorski et al. (2016), a parallel difference between Olympic and Non-Olympic competitions appears to be present for the 100 meter races (see their Figure 1). The authors even describe that “there was a small positive performance effect after adjusting for reaction time” (p. 412) for the swimming times of swimmers at positions 2 to 4 in Olympic 100 meter freestyle relays, which—nicely—replicates our findings from the 2012 study.⁶ The distribution of Olympic ($n = 144$) and Non-Olympic races ($n = 634$), thus, most likely biases the results of Skorski et al. (2016) towards effort losses during group work (i.e., slower swimming times in the relay than in the individual competition). The sample selected for the authors’ study appears not to be suited for showing that our findings may be unfounded, but rather analyzes another and a very specific population of relay competitions.

The third shortcoming: Skorski et al. select an arbitrary and rather small sample without providing a sufficient justification.

Above, we have illustrated that only the Olympic competition data from the Skorski et al. article is suited for showing that our findings are unfounded. The sample of Skorski et al. (2016) consists of $n = 144$ Olympic races and, thus, is rather small. It is not clear why the authors did not try to allocate a bigger sample (e.g., by including the data from the 1996 Olympics, which would also have been available on the website they used

⁵ For recent examples, see the nomination criteria for the 2017 World Championships by the German Swimming Association (www.dsv.de/schwimmen/nationalmannschaft/nominierungs-richtlinien/), the Austrian Swimming Association (<http://www.osv.or.at/schwimmen-mit-open-water/qualifikationsrichtlinien/>) or the US National Swimming Association (<https://www.usaswimming.org/resources-home/resource-topic/resource-subtopic>).

⁶ Note, however, that the exact numbers reported in Skorski et al. (2016) for these data are wrong or at least unclear: Although the depicted reaction-time-corrected effect sizes in their Figure 1 for the Olympic 100 meter competitions are clearly positive and appear to have an effect size d of about 0.50-0.60, the graphically depicted values are described in the figure caption as being either very small and positive (0.05) or even strongly negative (-1.06).

for their data collection [i.e., swimrankings.net]) and the authors do not provide a sufficient justification for the arbitrary selection of their sample.⁷ Assembling a certain sample size without a clear stopping rule (e.g., to collect all data that is available at the time of the study or to continue data collection until a stop criterion is reached, e.g., Frick, 1998) or without an a priori power analysis is problematic because it decreases the possibility to detect true effects (and increases the possibility of coming up with false positive effects; cf. Cohen, 1992).

The small sample size may have contributed to the absence of significant differences between relay and individual swimming performances—beyond the imprecise applied reaction-time correction (cf. the first shortcoming) and the predominant inclusion of national competition data (cf. the second shortcoming). With bigger samples and, thus, more statistical power, we have repeatedly found evidence for faster swimming times in the relay competition (see, for instance, the two studies in Hüffmeier et al., 2017).

The fourth shortcoming: Skorski et al. aggregate data over systematically different relay positions and may thereby obscure otherwise observable differences between relay and individual swimming times.

Skorski et al. (2016) aggregate the swimming times for swimmers at the relay positions 2 to 4 and only provide statistical results for this aggregate (i.e., no separate results per relay position are given). This is problematic because, in our studies, the size of the performance gains in the relay competition varied across positions. This information was available when Skorski et al. (2016) wrote their article (cf. Hüffmeier & Hertel, 2011; Hüffmeier et al., 2012)—but, admittedly, it became most evident in our most recent article that was published only later (Hüffmeier et al., 2017).

In the most recent article, we found only (very) small effort gains for positions 2 and 3, even in the condition that made effort gains in the relay most likely (i.e., high chances of winning a medal in the relay competition, high valence of the relay competition [e.g., at World championships]; see Hüffmeier et al., 2017). Specifically, we observed relay swimming times that were only 0.008 s (position 2) and 0.031 s (position 3) faster than the respective individual times—as compared to 0.202 s faster times at relay position 4. While

all these differences in swimming times were significantly different from zero in the Hüffmeier et al. (2017) study due to their large sample ($n = 738$ swimmers at positions 2 and 3 in 100 and 200 m races), studies that use smaller samples and aggregate swimming time differences across positions 2 to 4 might obscure otherwise observable faster swimming times especially at relay position 4.

Conclusion

In our discussion of the Skorski et al. (2016) article, we have identified the need to clarify the existing literature in one place and, more importantly, four shortcomings: While Skorski et al. suggest that controlling for reaction-time advantages in relays is a novel contribution of their study it is in fact common research practice. Turning to the first shortcoming, the authors applied imprecise correction factors. Second, they included data from national competitions and—due to the relative unattractiveness of national relay competitions—most likely biased their findings towards slow relay swimming times (cf. Hüffmeier et al., 2017; see also Karau et al., 2000). This data is not well suited for an unbiased empirical test of the hypothesis that relay swimming is not faster than individual swimming. Third, Skorski et al. selected an arbitrary and comparably small sample, thereby reducing the chance to find significant differences. Fourth and finally, they only provide aggregated data analyses despite well-documented differences between relay positions.

They, thus, obscure possibly otherwise observable differences between swimming times from relay and individual competitions especially at the last relay position. Given the multitude, weight, and association of these four shortcomings, we are convinced that the authors' findings are non-conclusive in view of their research question. Their data therefore do not question the validity of our findings and their conclusion—that our findings are “a myth”—is necessarily unfounded.

⁷ As far as we know, data availability in data bases starts with the 1996 Olympic Games.

Open Science Practices



This article earned the Open Data badge for making the data available through the Open Science Framework project page at <https://doi.org/10.17605/osf.io/rq4fv>. An editorial assistant independently confirmed that the code reproduced the results presented in the article. The editorial history can also be accessed from the OSF project page. The nature of the research question meant that there were no relevant research materials.

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